

December 1, 2017



Inter-University Research Institute Corporation /  
Research Organization of Information and Systems

**National Institute of Informatics**

# Bias-Variance Tradeoff

Data Mining 06 (データマイニング)

---

Mahito Sugiyama (杉山磨人)

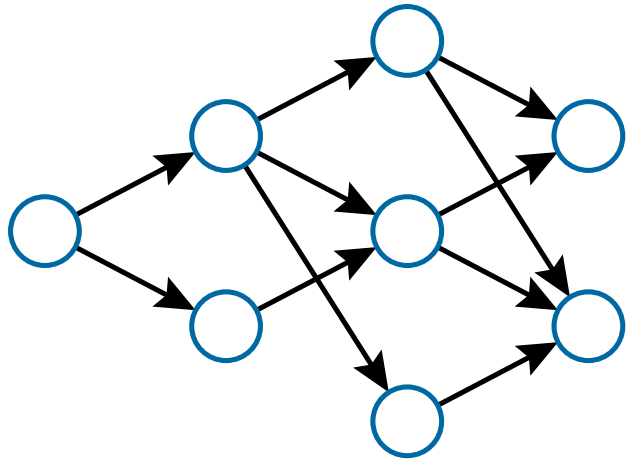
# Today's Outline

---

- Log-linear models on posets
  - A generalized formulation of Boltzmann machines
- Bias-variance tradeoff
- Fisher information & Cramér-Rao inequality

# Partially Ordered Set (Poset)

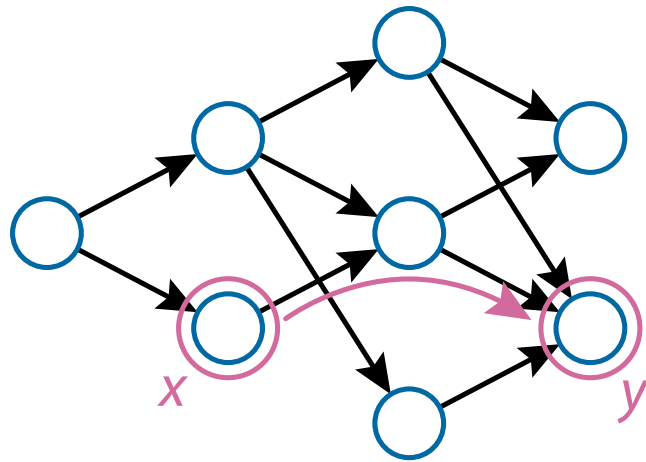
---



- Partially ordered set (**poset**)  $(S, \leq)$ 
  - $x \leq x$  (reflexivity)
  - $x \leq y, y \leq x \Rightarrow x = y$  (antisymmetry)
  - $x \leq y, y \leq z \Rightarrow x \leq z$  (transitivity)
    - We assume that  $S$  is finite and includes the least element (bottom)  $\perp \in S$

# Partially Ordered Set

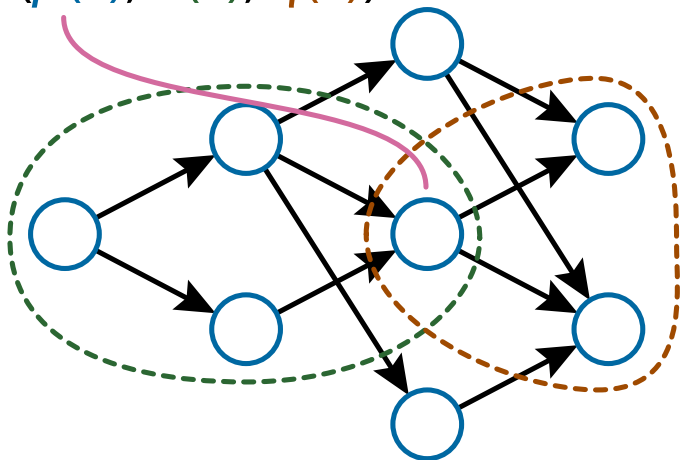
---



- Partially ordered set (**poset**)  $(S, \leq)$ 
  - (i)  $x \leq x$  (reflexivity)
  - (ii)  $x \leq y, y \leq x \Rightarrow x = y$  (antisymmetry)
  - (iii)  $x \leq y, y \leq z \Rightarrow x \leq z$  (transitivity)
    - We assume that  $S$  is finite and includes the least element (bottom)  $\perp \in S$
- Equivalent to a DAG
  - Each  $x \in S$  is a node
  - $x \leq y \iff y$  is reachable from  $x$

# Log-Linear Model on Poset

Each  $x \in S$  has a triple:  
 $(p(x), \theta(x), \eta(x))$



- A probability distribution  $p:S \rightarrow (0, 1)$   
s.t.  $\sum_{x \in S} p(x) = 1$
- We introduce  $\theta:S \rightarrow \mathbb{R}$  and  $\eta:S \rightarrow \mathbb{R}$  as
$$\log p(x) = \sum_{s \leq x} \theta(s)$$
$$\eta(x) = \sum_{s \geq x} p(s)$$
  - Parameter set  $B \subseteq S$
  - $\theta(s) = 0$  if  $s \notin B$

# Log-Linear Model on Powerset

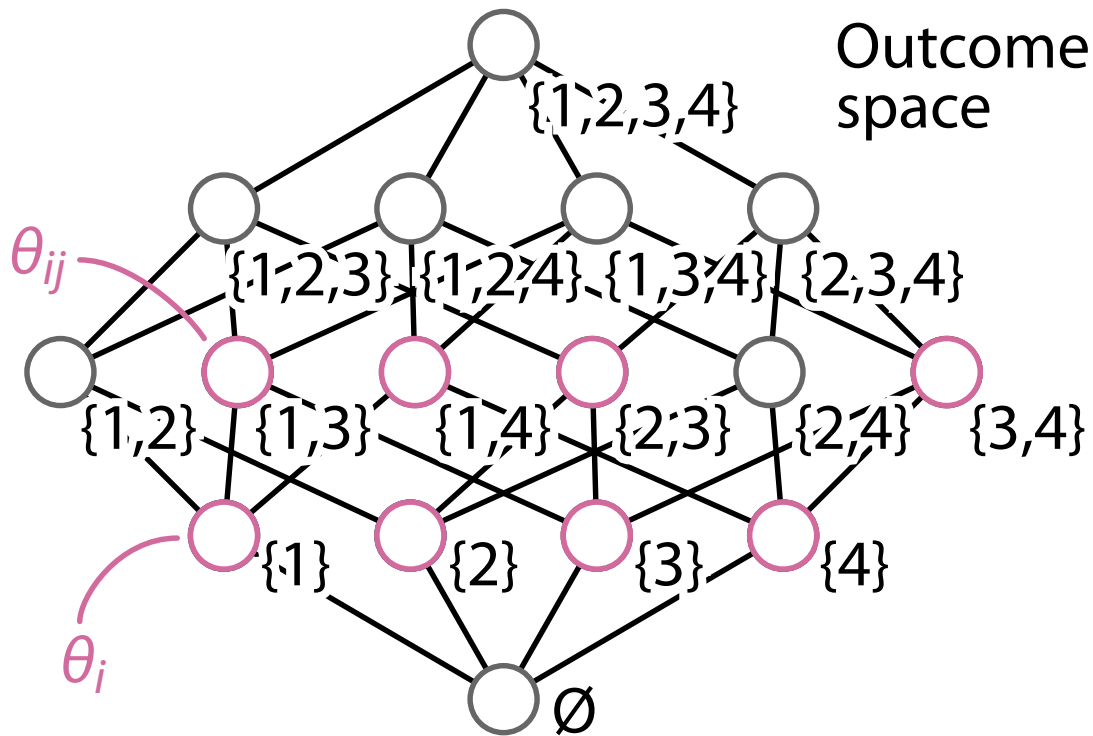
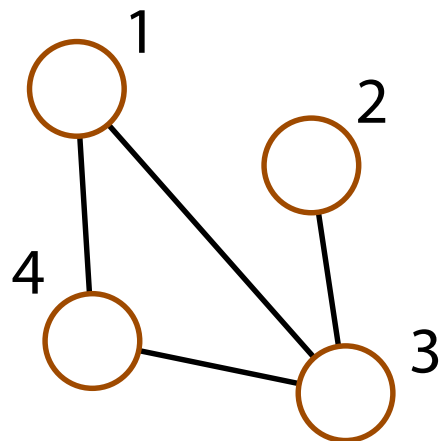
---

- Probability distribution over the power set  $2^V$  with  $V = \{1, 2, \dots, n\}$ 
  - $x \leq y \iff x \subseteq y, S = 2^V$
- Probability  $p(x)$  for each  $x \in 2^V$  is given as  $\log p(x) = \sum_{s \subseteq x} \theta(s)$ 
  - Parameter set  $B \subseteq 2^V$
  - $\theta(s) = 0$  if  $s \notin B$
- **MLE**: Find  $\theta(s)$  from a dataset  $D \subseteq 2^V$  for all  $s \in B$  s.t.  $\eta(s) = \hat{\eta}(s)$

$$\eta(s) = \sum_{x \supseteq s} p(x), \quad \hat{\eta}(s) = \frac{1}{|D|} \sum_{x \in D} \mathbf{1}[x \supseteq s] = |\{x \in D \mid x \supseteq s\}| / |D|$$

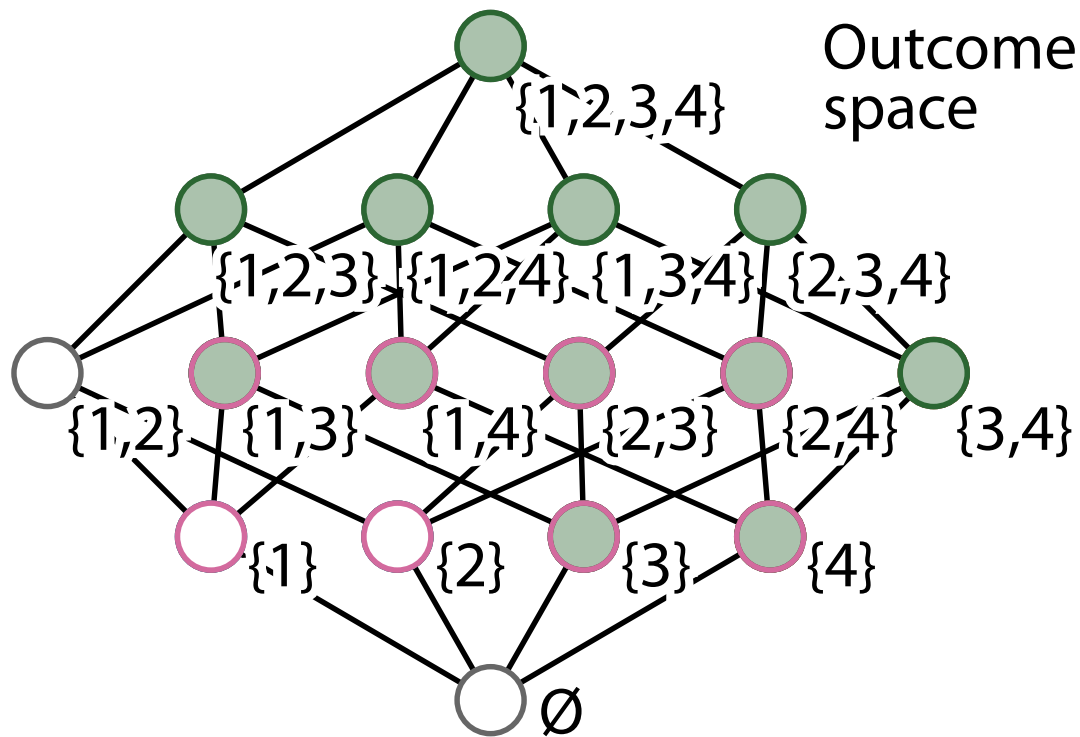
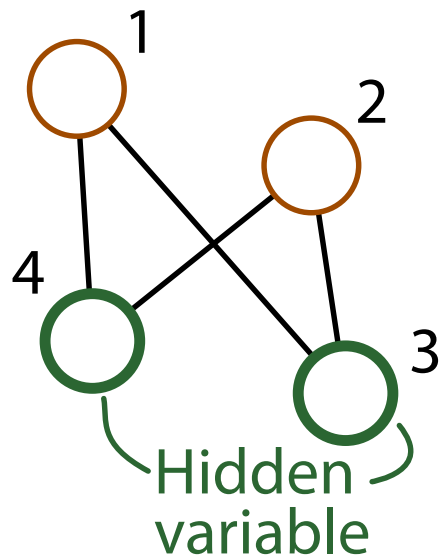
# Boltzmann Machines

Boltzmann Machine (BM)



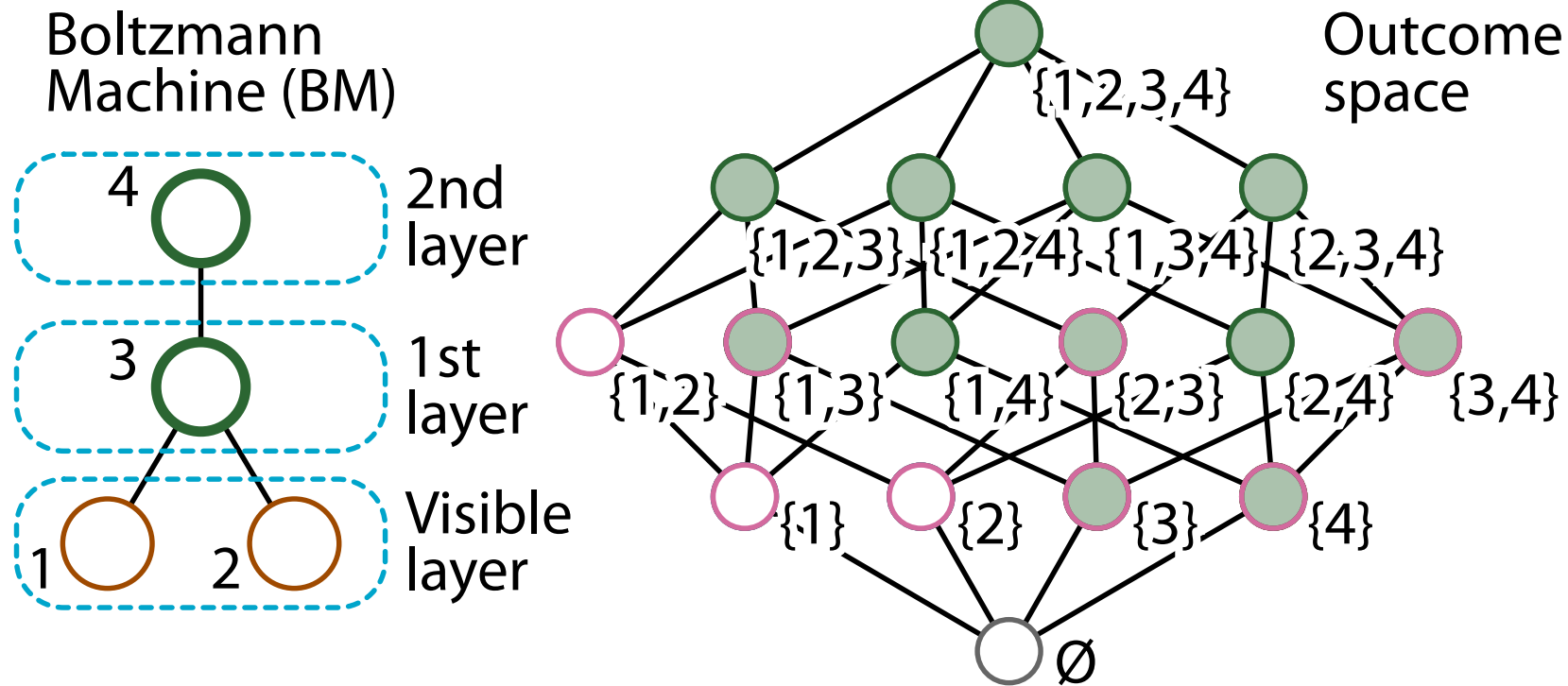
# Restricted Boltzmann Machines (RBMs)

Boltzmann Machine (BM)



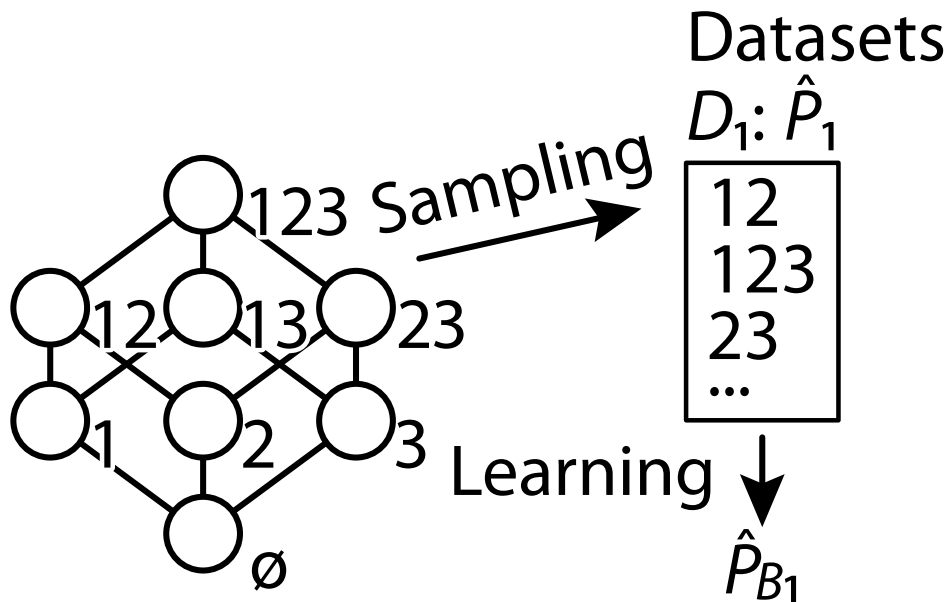


# Deep Boltzmann Machines (DBMs)



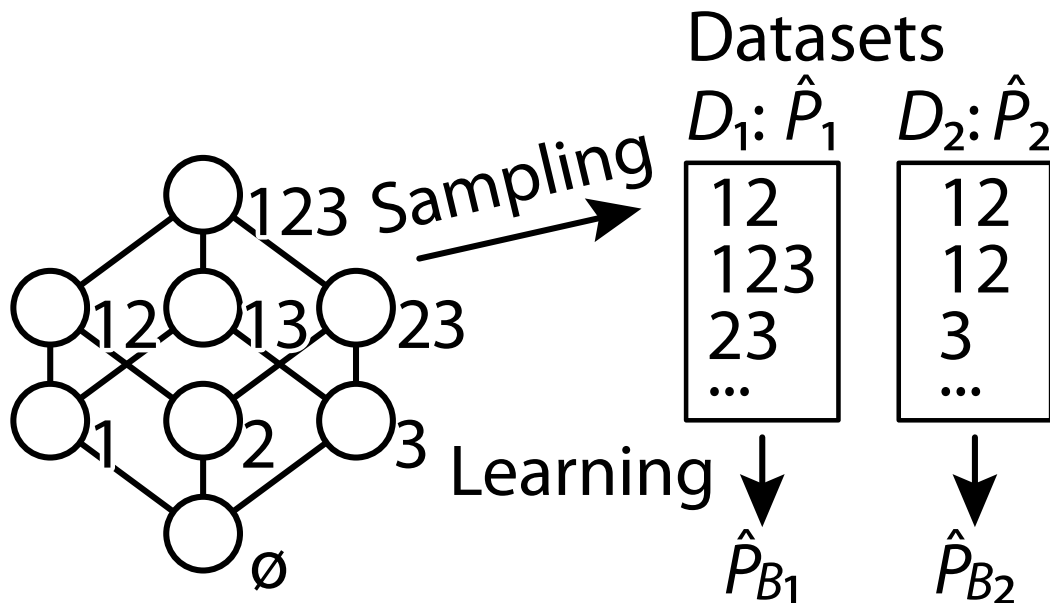
# Learning from Data

---



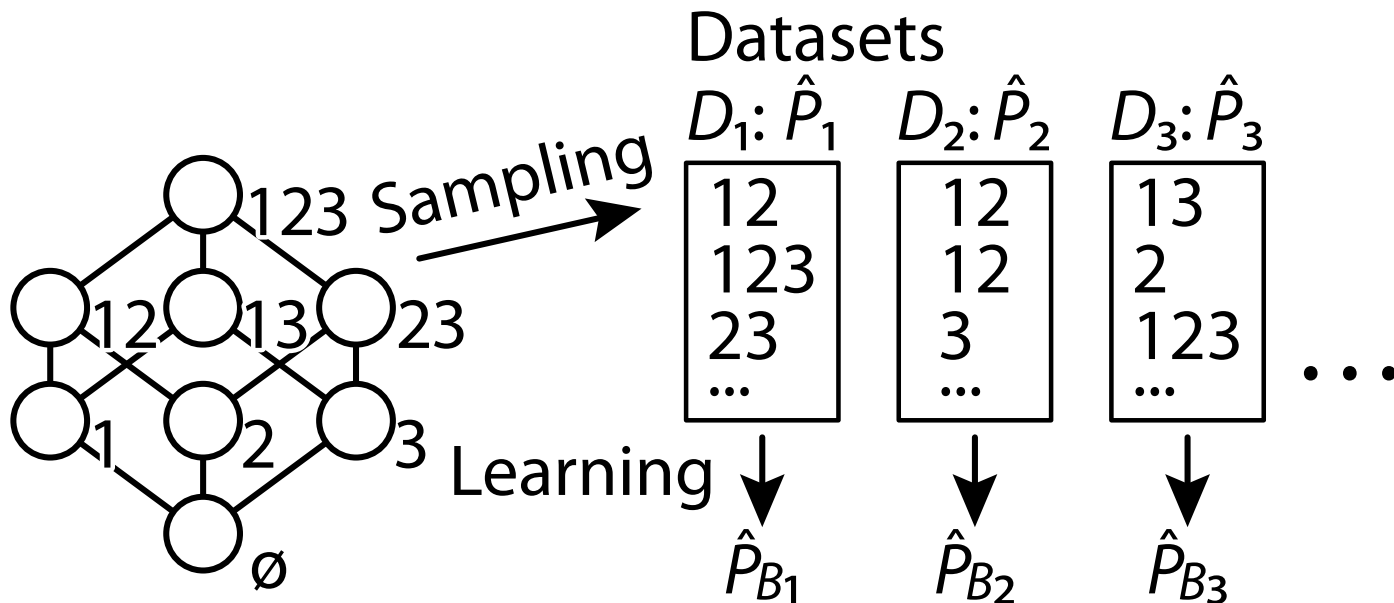
# Learning from Data

---



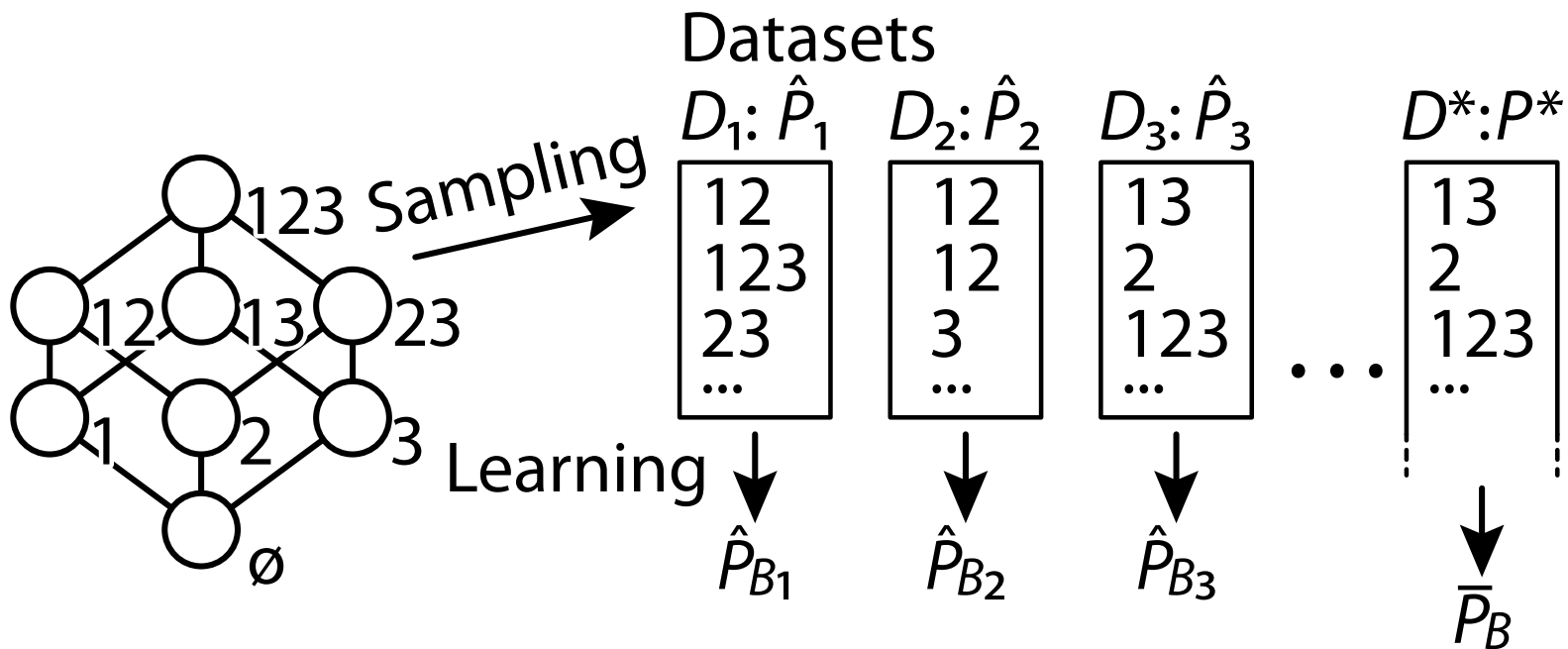
# Learning from Data

---

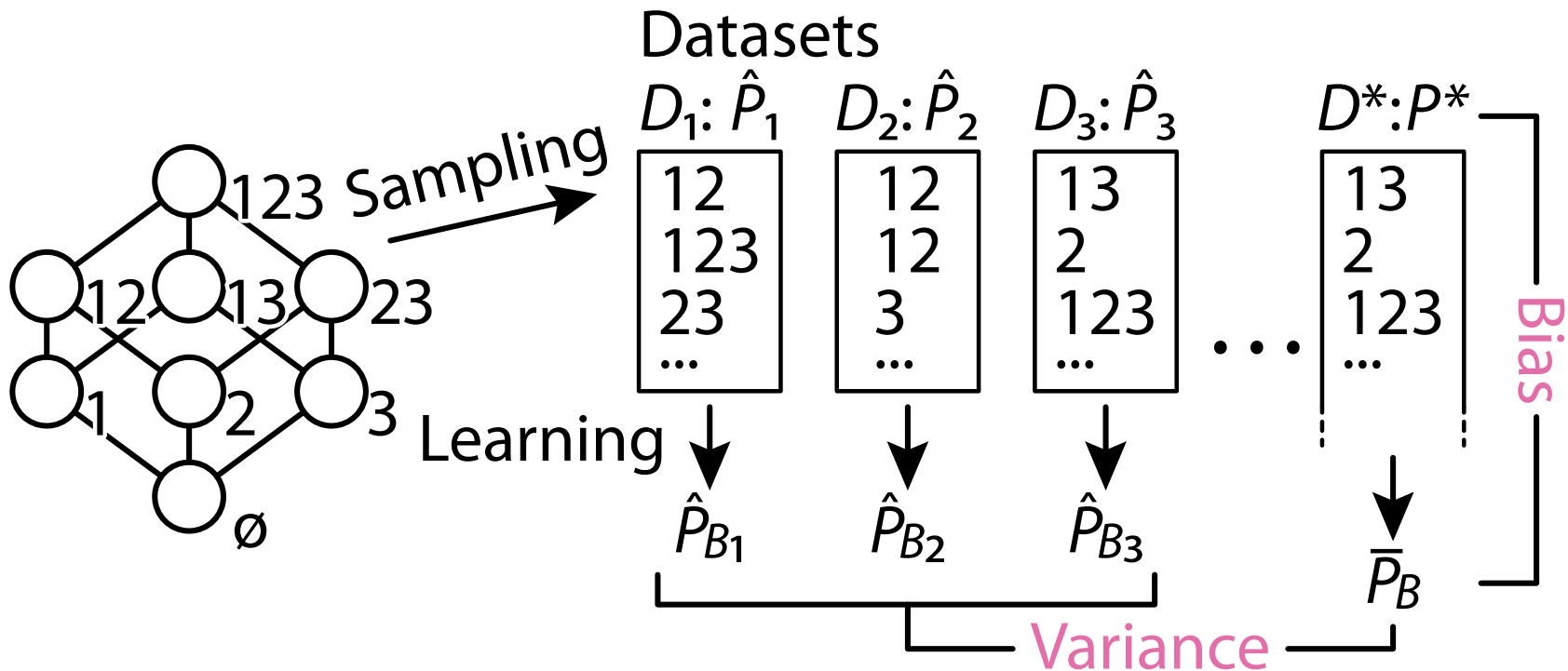


# Learning from Data

---



# Learning from Data



# Bias-Variance Tradeoff

---

- Bias =  $D_{\text{KL}}(P^*, \bar{P}_B)$
- Variance =  $\mathbf{E}[D_{\text{KL}}(\bar{P}_B, \hat{P}_B)]$
- If we include more parameters in  $B$ :
  - Bias will **decrease**
  - Variance will **increase**
- Two extreme cases:
  - If  $B = 2^V$ , then  $\hat{P}_B = \hat{P}$ , thus bias = 0 but variance will be large
  - If  $B = \emptyset$ ,  $\hat{P}_B$  is always the uniform distribution  $U$ , thus bias =  $D_{\text{KL}}(U, P^*)$  and variance = 0

# Bias-Variance Decomposition

---

- Decomposition of **MSE** (Mean Squared Error)

$$\begin{aligned}\mathbf{E}[(\hat{\theta} - \theta^*)^2] &= (\bar{\theta} - \theta^*)^2 + \mathbf{E}[(\hat{\theta} - \bar{\theta})^2] \\ &= \text{bias}^2(\hat{\theta}) + \text{var}[\hat{\theta}]\end{aligned}$$

$$\text{MSE} = \text{bias}^2 + \text{variance}$$

- $\theta^*$ : the true parameter
- $\hat{\theta}$ : the estimate
- $\bar{\theta}$ : the expected value of the estimate,  $\bar{\theta} = \mathbf{E}[\hat{\theta}]$   
(the estimate obtained from infinitely many data points)
- The expectation  $\mathbf{E}$  is about the true distribution  $p(D; \theta^*)$



# Example: Gaussian Mean Estimation

---

- Estimate the **mean** from  $N$  data points  $x_1, x_2, \dots, x_N$  sampled from a Gaussian distribution  $N(\theta^* = 1, \sigma^2)$

- Strategy 1: MLE

$$\text{bias} = \mathbf{E}[\hat{\theta}] - \theta^* = \mathbf{E}\left[\frac{1}{N} \sum_{i=1}^N x_i\right] - \theta^* = N\theta^*/N - \theta^* = 0$$

$$\text{variance} = \sigma^2/N$$

- Strategy 2: MAP estimate with a prior  $N(\theta_0, \sigma^2/\kappa_0)$

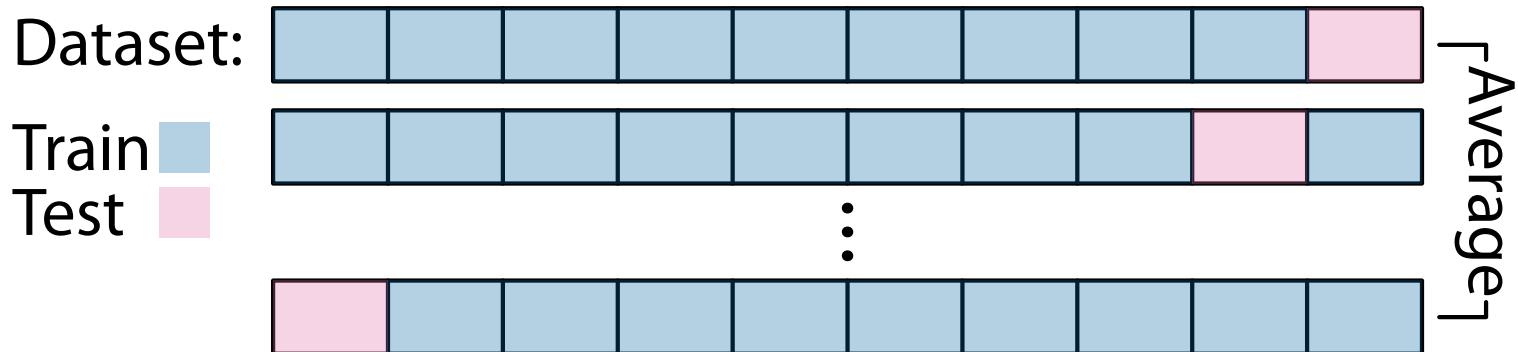
$$\text{bias} = \left(w\mathbf{E}[\hat{\theta}] + (1-w)\theta_0\right) - \theta^* = (1-w)(\theta_0 - \theta^*), \quad w = N/(N + \kappa_0),$$

$$\text{variance} = w^2 \sigma^2/N$$

# Practical Solution: Cross-Validation

---

- CV (Cross Validation) is the most convenient way to find the best parameter from data without seeing the true parameter
- K-fold cross-validation is typically used



# Fisher Information

---

- Let  $p(x; \xi)$  be a distribution with a parameter  $\xi$
- The **Fisher information**  $g(\xi)$  of  $\xi$  is

$$g(\xi) = \mathbf{E} \left[ \left( \frac{\partial}{\partial \xi} \log p(x; \xi) \right)^2 \right] = \sum_{x \in \mathcal{S}} p(x; \xi) \left( \frac{\partial}{\partial \xi} \log p(x; \xi) \right)^2$$

- If there are multiple parameters  $\boldsymbol{\xi} = (\xi_1, \xi_2, \dots, \xi_m)$ , the **Fisher information matrix** is a  $m \times m$  matrix  $G(\boldsymbol{\xi})$  given as

$$g(\boldsymbol{\xi})_{ij} = \mathbf{E} \left[ \frac{\partial}{\partial \xi_i} \log p(x; \boldsymbol{\xi}) \frac{\partial}{\partial \xi_j} \log p(x; \boldsymbol{\xi}) \right]$$

# Cramér-Rao Lower Bound

---

- Let  $\boldsymbol{\xi}$  be unbiased:  $\mathbf{E}[\hat{\boldsymbol{\xi}}] = \boldsymbol{\xi}^*$

- Cramér-Rao inequality:

$$E \geq \frac{1}{N} G(\boldsymbol{\xi})^{-1}$$

where  $E = (e_{ij})$ , each  $e_{ij} = \mathbf{E} \left[ (\hat{\xi}_i - \xi_i^*) (\hat{\xi}_j - \xi_j^*) \right]$

- $E$  coincides with the covariance matrix,  $e_{ii} = \mathbf{E} [ (\hat{\xi}_i - \xi_i^*)^2 ] = \text{var}(\hat{\xi}_i)$
  - $A > B$  if  $A - B$  is positive definite
    - $C$  is positive definite if  $\mathbf{x}^T C \mathbf{x} > 0$  for any non-zero  $\mathbf{x} \in \mathbb{R}^n$
- In MLE,  $E \rightarrow (1/N)G(\boldsymbol{\xi})^{-1}$  when  $N \rightarrow \infty$

# Example in Gaussian Mean Estimation

---

- Estimate the **mean** from  $N$  data points  $x_1, x_2, \dots, x_N$  sampled from a Gaussian distribution  $N(\theta^*, \sigma^2)$

- Fisher information:

$$g(\theta) = \frac{1}{\sigma^2}$$

- Cramér-Rao bound:

$$\text{var}[\hat{\theta}] \geq \frac{\sigma^2}{N}$$

- In this case,  $\text{var}[\hat{\theta}] = \sigma^2 / N$  always holds

# Model Selection by AIC

---

- The **AIC** (Akaike information criterion) is one of the most famous measure of the quality of statistical models

$$\text{AIC} = -2l(D) + 2k$$

- $l(D)$  is the maximized log-likelihood
  - $k$  is the number of parameters
- This cannot be directly used for log-linear models on posets as it is a hierarchical model with including higher-order associations
  - It is still under development...